## m3pm16l28.texLecture 28. 20.3.2012

To prove (\*):

$$\int_{2}^{x} \frac{\pi(x/u)}{\log u} du = \int_{2}^{x/2} \frac{\pi(x/u)}{\log u} du$$

(if u > x/2, x/u < 2, so  $\pi(x/u) = 0$ )

$$= x \int_2^{x/2} \frac{\pi(v)}{\log x - \log v} \frac{dv}{v^2}$$

 $\begin{array}{ll} (v := x/u; \ 2 \le u = x/v \le x/2, \ \text{so} \ 2 \le v \le x/2). \ \text{Choose} \ \epsilon > 0; \ \text{for} \\ v \ge V = V(\epsilon), \\ |\pi(v) - \frac{v}{\log w}| < \epsilon \frac{v}{\log w} \end{array}$ ľ

$$\pi(v) - \frac{v}{\log v}| < \epsilon \frac{v}{\log v}$$

by PNT. So for x > 2V,

$$\begin{split} |\int_{V}^{x/2} \frac{\pi(v)}{\log x - \log v} \frac{dv}{v^{2}} - \int_{V}^{x/2} \frac{v/\log v}{\log x - \log v} \frac{dv}{v^{2}}| &< \epsilon \int_{V}^{x/2} \frac{v/\log v}{\log x - \log v} \frac{dv}{v^{2}}, \\ \text{so } |\int_{2}^{x/2} \dots - \int_{2}^{x/2} \dots| &< \epsilon \int_{2}^{x/2} \dots + O(1/\log x), \\ \text{as} \\ \int_{2}^{V} \frac{v/\log v}{\log x - \log v} \frac{dv}{v^{2}} = O(1/\log x), \end{split}$$

etc. Since

$$\begin{split} \int_{2}^{x/2} \frac{v/\log v}{\log x - \log v} \frac{dv}{v^2} &= \int_{\log 2}^{\log x - \log 2} \frac{dw}{w(\log x - w)} \qquad (w := \log v) \\ &= \frac{1}{\log x} \int_{\log 2}^{\log x - \log 2} \left(\frac{1}{w} + \frac{1}{\log x - w}\right) dw \qquad (\text{partial fractions}) \\ &= \frac{1}{\log x} (\log(\log x - \log 2) - \log\log 2 - \log\log 2 + \log(\log x - \log 2)) \sim \frac{2\log\log x}{\log x}, \\ \text{this gives (*).} \end{split}$$

By the Lemma,  $\pi_2(x) \sim x \log \log x / \log x$ , proving the case k = 2. //

As each n has at least one prime factor, it is better to work with k + 1rather than k. Writing

$$\lambda := \log \log x$$

(so  $\lambda \to \infty$  as  $x \to \infty$  – though *extremely slowly*):

$$\frac{1}{x}\pi_{k+1}(x) \sim \frac{(\log\log x)^k}{k!\log x} = \frac{e^{-\lambda}\lambda^k}{k!} \quad (k=0,1,2,\ldots) \qquad (\lambda,x\to\infty).$$

Now  $\{e^{-\lambda}\lambda^k/k!: k = 0, 1, 2, ...\}$  forms the *Poisson distribution*  $P(\lambda)$  of Probability Theory, with parameter  $\lambda$  (mean  $\lambda$ , variance  $\lambda$ ). So:

**Theorem (Landau)**. The proportion of primes  $\leq x$  with k+1 distinct prime factors is asymptotically Poisson distributed with parameter  $\lambda := \log \log x$ .

The Poisson distribution is "the signature of randomness", in the *discrete* setting (as here). So this suggests that, *in some sense, the primes are randomly distributed* (hence 'Primes play a game of chance' – see III.10.1 below). This is very surprising: in the ordinary sense, nothing could be less random, or more deterministic, or "God-given", than the primes.

Recall the prime divisor functions of II.8:  $\omega(n)$  is the number of distinct prime divisors of n,  $\Omega(n)$  is the number of prime divisors of n counted with multiplicity. It turns out that, as in II.8,  $\omega$  and  $\Omega$  behave similarly here. So we may rephrase Landau's theorem above as saying that both proportions  $\omega(n)/n$ ,  $\Omega(n)/n$  are asymptotically Poisson distributed with parameter  $\lambda := \log \log n$  (recall that  $\log \log played$  a key role in II.8 also). Using  $X \sim F$ as the usual probabilistic shorthand for "the random variable X has the distribution (function) F", we have

## Theorem (Landau's Poisson PNT, 1900).

$$\omega(n)/n \sim P(\log \log n), \qquad \Omega(n)/n \sim P(\log \log n).$$

With some loss of information (the constants  $C_1, C_2$  and the error terms  $O(x/\log x)$ , we may summarise the Theorem of II.8 for comparison. Using  $\sim$  now (with a number after it, not a distribution) to denote "is asymptotic to", one has

Theorem (Hardy and Ramanujan, 1917).

$$\omega(n)/n \sim \log \log n, \qquad \Omega(n) \sim \log \log n.$$