

**Lecture 17 17.10.2015**

To find the (perfect-hedge) strategy for replicating this explicitly: write

$$c(n, x) := \sum_{j=0}^{N-n} \binom{N-n}{j} p^{*j} (1-p^*)^{N-n-j} (x(1+a)^j (1+b)^{N-n-j} - K)_+.$$

Then  $c(n, x)$  is the undiscounted  $P^*$ -expectation of the call at time  $n$  given that  $S_n = x$ . This must be the value of the portfolio at time  $n$  if the strategy  $H = (H_n)$  replicates the claim:

$$H_n^0(1+r)^n + H_n S_n = c(n, S_n)$$

(here by previsibility  $H_n^0$  and  $H_n$  are both functions of  $S_0, \dots, S_{n-1}$  only). Now  $S_n = S_{n-1}T_n = S_{n-1}(1+a)$  or  $S_{n-1}(1+b)$ , so:

$$\begin{aligned} H_n^0(1+r)^n + H_n S_{n-1}(1+a) &= c(n, S_{n-1}(1+a)) \\ H_n^0(1+r)^n + H_n S_{n-1}(1+b) &= c(n, S_{n-1}(1+b)). \end{aligned}$$

Subtract:

$$H_n S_{n-1}(b-a) = c(n, S_{n-1}(1+b)) - c(n, S_{n-1}(1+a)).$$

So  $H_n$  in fact depends only on  $S_{n-1}$ ,  $H_n = H_n(S_{n-1})$  (by previsibility), and

**Proposition.** The perfect hedging strategy  $H_n$  replicating the European call option above is given by

$$H_n = H_n(S_{n-1}) = \frac{c(n, S_{n-1}(1+b)) - c(n, S_{n-1}(1+a))}{S_{n-1}(b-a)}.$$

Notice that the numerator is the difference of two values of  $c(n, x)$  with the larger value of  $x$  in the first term (recall  $b > a$ ). When the payoff function  $c(n, x)$  is an increasing function of  $x$ , as for the European call option considered here, this is non-negative. In this case, the Proposition gives  $H_n \geq 0$ : the replicating strategy does not involve short-selling. We record this as:

**Corollary.** When the payoff function is a non-decreasing function of the final asset price  $S_N$ , the perfect-hedging strategy replicating the claim does not involve short-selling of the risky asset.

### §6. Continuous-Time Limit of the Binomial Model.

Suppose now that we wish to price an option in continuous time with initial stock price  $S_0$ , strike price  $K$  and expiry  $T$ . We can use the work above to give a discrete-time approximation, where  $N \rightarrow \infty$ . Given  $R \geq 0$ , the instantaneous interest rate in continuous time, define  $r$  by

$$r := RT/N : \quad e^{RT} = \lim_{N \rightarrow \infty} \left(1 + \frac{RT}{N}\right)^N = \lim_{N \rightarrow \infty} (1 + r)^N.$$

Here  $r$ , which tends to zero as  $N \rightarrow \infty$ , represents the interest rate in discrete time for the approximating binomial model.

For  $\sigma > 0$  fixed ( $\sigma^2$  plays the role of a variance, corresponding in continuous time to the *volatility* of the stock – below), define  $a, b$  ( $\rightarrow 0$  as  $N \rightarrow \infty$ ) by

$$\log((1+a)/(1+r)) = -\sigma/\sqrt{N}, \quad \log((1+b)/(1+r)) = \sigma/\sqrt{N}.$$

We now have a sequence of binomial models, for each of which we can price options as in §5. We shall show that the pricing formula converges as  $N \rightarrow \infty$  to a limit. This is the famous *Black-Scholes formula*, the central result of the course. We shall meet it later, and re-derive it, in *continuous time*, its natural setting, in Ch. VI; see also e.g. [BK], 4.6.2.

**Lemma.** Let  $(X_j^N)_{j=1}^N$  be iid with mean  $\mu_N$  satisfying

$$N\mu_N \rightarrow \mu \quad (N \rightarrow \infty)$$

and variance  $\sigma^2(1+o(1))/N$ . If  $Y_N := \sum_1^N X_j^N$ , then  $Y_N$  converges in distribution to normality:

$$Y_N \rightarrow Y = N(\mu, \sigma^2) \quad (N \rightarrow \infty).$$

*Proof.* Use characteristic functions (CFs): since  $Y_N$  has mean and variance as given, it also has second moment  $\sigma^2(1+o(1))/N$ , so has CF

$$\begin{aligned} \phi_N(u) &:= E \exp\{iuY_N\} = \prod_1^N E \exp\{iuX_j^N\} = [E \exp\{iuX_1^N\}]^N \\ &= \left(1 + \frac{i u \mu}{N} - \frac{1}{2} \frac{\sigma^2 u^2}{N} + o\left(\frac{1}{N}\right)\right)^N \rightarrow \exp\left\{i u \mu - \frac{1}{2} \sigma^2 u^2\right\} \quad (N \rightarrow \infty), \end{aligned}$$

the CF of the normal law  $N(\mu, \sigma^2)$ . Convergence of CFs implies convergence in distribution by Lévy's continuity theorem for CFs ([W], §18.1). //

We can apply this to pricing the call option above:

$$\begin{aligned} C_0^{(N)} &= \left(1 + \frac{RT}{N}\right)^{-N} E^*[(S_0 \Pi_1^N T_n - K)_+] \\ &= E^*[(S_0 \exp\{Y_N\} - (1 + \frac{RT}{N})^{-N} K)_+], \end{aligned} \quad (1)$$

where

$$Y_N := \sum_1^N \log(T_n/(1+r)).$$

Since  $T_n = T_n^N$  above takes values  $1+b, 1+a$ ,  $X_n^N := \log(T_n^N/(1+r))$  takes values  $\log((1+b)/(1+r)), \log((1+a)/(1+r)) = \pm\sigma/\sqrt{N}$  (so has second moment  $\sigma^2/N$ ). Its mean is

$$\mu_N := \log\left(\frac{1+b}{1+r}\right)(1-p^*) + \log\left(\frac{1+a}{1+r}\right)p^* = \frac{\sigma}{\sqrt{N}}(1-p^*) - \frac{\sigma}{\sqrt{N}}p^* = (1-2p^*)\sigma/\sqrt{N}$$

(we shall see below that  $1-2p^* = O(1/\sqrt{N})$ , so the Lemma will apply). Now (recall  $r = RT/N = O(1/N)$ )

$$a = (1+r)e^{-\sigma/\sqrt{N}} - 1, \quad b = (1+r)e^{\sigma/\sqrt{N}} - 1,$$

so  $a, b, r \rightarrow 0$  as  $N \rightarrow \infty$ , and

$$\begin{aligned} 1 - 2p^* &= 1 - 2\frac{(b-r)}{(b-a)} = 1 - 2\frac{[(1+r)e^{\sigma/\sqrt{N}} - 1 - r]}{[(1+r)(e^{\sigma/\sqrt{N}} - e^{-\sigma/\sqrt{N}})]} \\ &= 1 - 2\frac{[e^{\sigma/\sqrt{N}} - 1]}{[e^{\sigma/\sqrt{N}} - e^{-\sigma/\sqrt{N}}]}. \end{aligned}$$

Now expand the two  $[\dots]$  terms above by Taylor's theorem: they give

$$\frac{\sigma}{\sqrt{N}}\left(1 + \frac{1}{2}\frac{\sigma}{\sqrt{N}} + \dots\right), \quad \frac{2\sigma}{\sqrt{N}}\left(1 + \frac{\sigma^2}{6N} + \dots\right).$$

So, cancelling  $\sigma/\sqrt{N}$ ,

$$1 - 2p^* = 1 - \frac{2\left(1 + \frac{1}{2}\frac{\sigma}{\sqrt{N}} + \dots\right)}{2\left(1 + \frac{\sigma^2}{6N} + \dots\right)} = -\frac{1}{2}\frac{\sigma}{\sqrt{N}} + O(1/N):$$

$$N\mu_N = N \cdot \frac{\sigma}{\sqrt{N}} \cdot \left(-\frac{1}{2}\frac{\sigma}{\sqrt{N}} + O(1/N)\right) \rightarrow \mu := -\frac{1}{2}\sigma^2 \quad (N \rightarrow \infty).$$

We now need to change notation:

(i) We replace the variance  $\sigma^2$  above by  $\sigma^2 T$ . So  $\sigma^2$  is the *variance per unit time* (which is more suited to the work of Ch. V, VI in continuous time); the standard deviation (SD)  $\sigma$  is called the *volatility*. It measures the variability of the stock, so its riskiness, or its sensitivity to new information.

(ii) We replace  $R$  in the above by  $r$ . This is the standard notation for the riskless interest rate in continuous time, to which we are now moving.

As usual, we write the standard normal density function as  $\phi$  and distribution function as  $\Phi$ :

$$\phi(x) := \frac{e^{-\frac{1}{2}x^2}}{\sqrt{2\pi}}, \quad \Phi(x) := \int_{-\infty}^x \phi(u) du = \int_{-\infty}^x \frac{e^{-\frac{1}{2}u^2}}{\sqrt{2\pi}} du.$$

Note that as  $\phi$  is even, the left and right tails of  $\Phi$  are equal:

$$\phi(x) = \phi(-x), \quad \text{so} \quad \int_{-\infty}^{-x} \phi(u) du = \int_x^{\infty} \phi(u) du : \quad \Phi(-x) = 1 - \Phi(x).$$

**Theorem (Black-Scholes formula (for calls), 1973).** The price of the European call option is

$$c_t = S_t \Phi(d_+) - K e^{-r(T-t)} \Phi(d_-), \quad (BS)$$

where  $S_t$  is the stock price at time  $t \in [0, T]$ ,  $K$  is the strike price,  $r$  is the riskless interest rate,  $\sigma$  is the volatility and

$$d_{\pm} := [\log(S/K) + (r \pm \frac{1}{2}\sigma^2)(T-t)] / \sigma \sqrt{T-t} : \quad d_+ = d_- + \sigma \sqrt{T-t}.$$

For completeness, we state the corresponding Black-Scholes formula for puts. The proofs of the two results are closely analogous, and one can derive either from the other by put-call parity.

**Theorem (Black-Scholes formula for puts, 1973).** The price of the corresponding put option is

$$p_t = K e^{-r(T-t)} \Phi(-d_-) - S_t \Phi(-d_+). \quad (BS - p)$$

The Black-Scholes formula is not perfect – indeed, Fischer Black himself famously wrote a paper called *The holes in Black-Scholes*. But it is very useful, as a benchmark and first approximation.