

m3f22117tex

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To find the (perfect-hedge) strategy for replicating this explicitly: write

$$c(n, x) := \sum_{j=0}^{N-n} \binom{N-n}{j} p^{*j} (1-p^*)^{N-n-j} (x(1+a)^j (1+b)^{N-n-j} - K)_+.$$

Then $c(n, x)$ is the undiscounted P^* -expectation of the call at time n given that $S_n = x$. This must be the value of the portfolio at time n if the strategy $H = (H_n)$ replicates the claim:

$$H_n^0(1+r)^n + H_n S_n = c(n, S_n)$$

(here by previsibility H_n^0 and H_n are both functions of S_0, \dots, S_{n-1} only). Now $S_n = S_{n-1}T_n = S_{n-1}(1+a)$ or $S_{n-1}(1+b)$, so:

$$\begin{aligned} H_n^0(1+r)^n + H_n S_{n-1}(1+a) &= c(n, S_{n-1}(1+a)) \\ H_n^0(1+r)^n + H_n S_{n-1}(1+b) &= c(n, S_{n-1}(1+b)). \end{aligned}$$

Subtract:

$$H_n S_{n-1}(b-a) = c(n, S_{n-1}(1+b)) - c(n, S_{n-1}(1+a)).$$

So H_n in fact depends only on S_{n-1} , $H_n = H_n(S_{n-1})$ (by previsibility), and

Proposition. The perfect hedging strategy H_n replicating the European call option above is given by

$$H_n = H_n(S_{n-1}) = \frac{c(n, S_{n-1}(1+b)) - c(n, S_{n-1}(1+a))}{S_{n-1}(b-a)}.$$

Notice that the numerator is the difference of two values of $c(n, x)$ with the larger value of x in the first term (recall $b > a$). When the payoff function $c(n, x)$ is an increasing function of x , as for the European call option considered here, this is non-negative. In this case, the Proposition gives $H_n \geq 0$: the replicating strategy does not involve short-selling. We record this as:

Corollary. When the payoff function is a non-decreasing function of the final asset price S_N , the perfect-hedging strategy replicating the claim does not involve short-selling of the risky asset.

§6. Continuous-Time Limit of the Binomial Model.

Suppose now that we wish to price an option in continuous time with initial stock price S_0 , strike price K and expiry T . We can use the work above to give a discrete-time approximation, where $N \rightarrow \infty$. We write (temporarily) $\rho \geq 0$ for the instantaneous interest rate in continuous time, and define (again temporarily) r by

$$r := \rho T/N : \quad e^{\rho T} = \lim_{N \rightarrow \infty} \left(1 + \frac{\rho T}{N}\right)^N = \lim_{N \rightarrow \infty} (1 + r)^N.$$

Here r , which tends to zero as $N \rightarrow \infty$, represents the interest rate in discrete time for the approximating binomial model.

For $\sigma > 0$ fixed (σ^2 plays the role of a variance, corresponding in continuous time to the *volatility* of the stock – below), define a, b ($\rightarrow 0$ as $N \rightarrow \infty$) by

$$\log((1+a)/(1+r)) = -\sigma/\sqrt{N}, \quad \log((1+b)/(1+r)) = \sigma/\sqrt{N}.$$

We now have a sequence of binomial models, for each of which we can price options as in §5. We shall show that the pricing formula converges as $N \rightarrow \infty$ to a limit. This is the famous *Black-Scholes formula*, the central result of the course. We shall meet it later, and re-derive it, in *continuous time*, its natural setting, in Ch. VI; see also e.g. [BK], 4.6.2.

Lemma. Let $(X_j^N)_{j=1}^N$ be iid with mean μ_N satisfying

$$N\mu_N \rightarrow \mu \quad (N \rightarrow \infty)$$

and variance $\sigma^2(1+o(1))/N$. If $Y_N := \sum_1^N X_j^N$, then Y_N converges in distribution to normality:

$$Y_N \rightarrow Y = N(\mu, \sigma) \quad (N \rightarrow \infty).$$

Proof. Use characteristic functions (CFs): since Y_N has mean and variance as given, it also has second moment $\sigma^2(1+o(1))/N$, so has CF

$$\begin{aligned} \phi_N(u) &:= E \exp\{iuY_N\} = \Pi_1^N E \exp\{iuX_j^N\} = [E \exp\{iuX_1^N\}]^N \\ &= \left(1 + \frac{iu\mu}{N} - \frac{1}{2} \frac{\sigma^2 u^2}{N} + o\left(\frac{1}{N}\right)\right)^N \rightarrow \exp\left\{iu\mu - \frac{1}{2} \sigma^2 u^2\right\} \quad (N \rightarrow \infty), \end{aligned}$$

the CF of the normal law $N(\mu, \sigma)$. Convergence of CFs implies convergence in distribution by Lévy's continuity theorem for CFs ([W], §18.1). //

We can apply this to pricing the call option above:

$$\begin{aligned} C_0^{(N)} &= \left(1 + \frac{\rho T}{N}\right)^{-N} E^*[(S_0 \Pi_1^N T_n - K)_+] \\ &= E^*[(S_0 \exp\{Y_N\} - (1 + \frac{\rho T}{N})^{-N} K)_+], \end{aligned} \quad (1)$$

where

$$Y_N := \sum_1^N \log(T_n/(1+r)).$$

Since $T_n = T_n^N$ above takes values $1+b, 1+a$, $X_n^N := \log(T_n^N/(1+r))$ takes values $\log((1+b)/(1+r)), \log((1+a)/(1+r)) = \pm\sigma/\sqrt{N}$ (so has second moment σ^2/N). Its mean is

$$\mu_N := \log\left(\frac{1+b}{1+r}\right)(1-p^*) + \log\left(\frac{1+a}{1+r}\right)p^* = \frac{\sigma}{\sqrt{N}}(1-p^*) - \frac{\sigma}{\sqrt{N}}p^* = (1-2p^*)\sigma/\sqrt{N}$$

(we shall see below that $1-2p^* = O(1/\sqrt{N})$, so the Lemma will apply). Now (recall $r = \rho T/N = O(1/N)$)

$$a = (1+r)e^{-\sigma/\sqrt{N}} - 1, \quad b = (1+r)e^{\sigma/\sqrt{N}} - 1,$$

so $a, b, r \rightarrow 0$ as $N \rightarrow \infty$, and

$$\begin{aligned} 1 - 2p^* &= 1 - 2 \frac{(b-r)}{(b-a)} = 1 - 2 \frac{[(1+r)e^{\sigma/\sqrt{N}} - 1 - r]}{[(1+r)(e^{\sigma/\sqrt{N}} - e^{-\sigma/\sqrt{N}})]} \\ &= 1 - 2 \frac{[e^{\sigma/\sqrt{N}} - 1]}{[e^{\sigma/\sqrt{N}} - e^{-\sigma/\sqrt{N}}]}. \end{aligned}$$

Now expand the two $[\dots]$ terms above by Taylor's theorem: they give

$$\frac{\sigma}{\sqrt{N}}\left(1 + \frac{1}{2} \frac{\sigma}{\sqrt{N}} + \dots\right), \quad \frac{2\sigma}{\sqrt{N}}\left(1 + \frac{\sigma^2}{6N} + \dots\right).$$

So, cancelling σ/\sqrt{N} ,

$$1 - 2p^* = 1 - \frac{2\left(1 + \frac{1}{2} \frac{\sigma}{\sqrt{N}} + \dots\right)}{2\left(1 + \frac{\sigma^2}{6N} + \dots\right)} = -\frac{1}{2} \frac{\sigma}{\sqrt{N}} + O(1/N):$$

$$N\mu_N = N \cdot \frac{\sigma}{\sqrt{N}} \cdot \left(-\frac{1}{2} \frac{\sigma}{\sqrt{N}} + O(1/N)\right) \rightarrow \mu := -\frac{1}{2}\sigma^2 \quad (N \rightarrow \infty).$$

We now need to change notation:

(i) We replace the variance σ^2 above by $\sigma^2 T$. So σ^2 is the *variance per unit time* (which is more suited to the work of Ch. V, VI in continuous time); the standard deviation (SD) σ is called the *volatility*. It measures the variability of the stock, so its riskiness, or its sensitivity to new information.

(ii) We replace ρ in the above by r . This is the standard notation for the riskless interest rate in continuous time, to which we are now moving.

As usual, we write the standard normal density function as ϕ and distribution function as Φ :

$$\phi(x) := \frac{e^{-\frac{1}{2}x^2}}{\sqrt{2\pi}}, \quad \Phi(x) := \int_{-\infty}^x \phi(u) du = \int_{-\infty}^x \frac{e^{-\frac{1}{2}u^2}}{\sqrt{2\pi}} du.$$

Note that as ϕ is even, the left and right tails of Φ are equal:

$$\phi(x) = \phi(-x), \quad \text{so} \quad \int_{-\infty}^{-x} \phi(u) du = \int_x^{\infty} \phi(u) du : \quad \Phi(-x) = 1 - \Phi(x).$$

Theorem (Black-Scholes formula (for calls), 1973). The price of the European call option is

$$c_t = S_t \Phi(d_+) - K e^{-r(T-t)} \Phi(d_-), \quad (BS)$$

where S_t is the stock price at time $t \in [0, T]$, K is the strike price, r is the riskless interest rate, σ is the volatility and

$$d_{\pm} := [\log(S/K) + (r \pm \frac{1}{2}\sigma^2)(T-t)] / \sigma \sqrt{T-t} : \quad d_+ = d_- + \sigma \sqrt{T-t}.$$

For completeness, we state the corresponding Black-Scholes formula for puts. The proofs of the two results are closely analogous, and one can derive either from the other by put-call parity.

Theorem (Black-Scholes formula for puts, 1973). The price of the corresponding put option is

$$p_t = K e^{-r(T-t)} \Phi(-d_-) - S_t \Phi(-d_+). \quad (BS - p)$$

The Black-Scholes formula is not perfect – indeed, Fischer Black himself famously wrote a paper called *The holes in Black-Scholes*. But it is very useful, as a benchmark and first approximation.