

London Taught Course Centre: Measure-Theoretic Probability
Solutions to Examination, 2010

Q1. (i)

$$\begin{aligned}
 \psi(t) = E[e^{itY}] &= E[\exp\{it(X_1 + \dots + X_N)\}] \\
 &= \sum_n E[\exp\{it(X_1 + \dots + X_N)\} | N = n] \cdot P(N = n) \\
 &= \sum_n e^{-\lambda} \lambda^n / n! \cdot E[\exp\{it(X_1 + \dots + X_n)\}] \\
 &= \sum_n e^{-\lambda} \lambda^n / n! \cdot (E[\exp\{itX_1\}])^n \\
 &= \sum_n e^{-\lambda} \lambda^n / n! \cdot \phi(t)^n \\
 &= \exp\{-\lambda(1 - \phi(t))\}.
 \end{aligned}$$

Differentiate:

$$\begin{aligned}
 \psi'(t) &= \psi(t) \cdot \lambda \phi'(t), \\
 \psi''(t) &= \psi'(t) \cdot \lambda \phi'(t) + \psi(t) \cdot \lambda \phi''(t).
 \end{aligned}$$

As $\phi(t) = E[e^{itX}]$, $\phi'(t) = E[iXe^{itX}]$, $\phi''(t) = E[-X^2e^{itX}]$. So $(\phi(0) = 1$ and) $\phi'(0) = i\mu$, $\phi''(0) = -E[X^2]$,

$$\psi'(0) = \lambda \phi'(0) = \lambda i\mu,$$

and as also $\psi'(0) = iEY$, this gives $EY = \lambda\mu$. Similarly,

$$\psi''(0) = i\lambda\mu \cdot i\lambda\mu + \lambda\phi''(0) = -\lambda^2\mu^2 - \lambda E[X^2],$$

and also $(\psi(0) = 1, \psi'(0) = i\lambda\mu$ and) $\psi''(0) = -E[Y^2]$. So

$$\text{var } Y = E[Y^2] - [EY]^2 = \lambda^2\mu^2 + \lambda E[X^2] - \lambda^2\mu^2 = \lambda E[X^2].$$

(ii) Given N , $Y = X_1 + \dots + X_N$ has mean $NEX = N\mu$ and variance $N \text{ var } X = N\sigma^2$. As N is Poisson with parameter λ , N has mean λ and variance λ . So by the Conditional Mean Formula,

$$EY = E[E(Y|N)] = E[N\mu] = \lambda\mu.$$

By the Conditional Variance Formula,

$$\begin{aligned}
 \text{var } Y &= E[\text{var}(Y|N)] + \text{var } E[Y|N] = E[N \text{ var } X] + \text{var}[N EX] \\
 &= EN \cdot \text{var } X + \text{var } N \cdot (EX)^2 = \lambda[E(X^2) - (EX)^2] + \lambda \cdot (EX)^2 = \lambda E[X^2].
 \end{aligned}$$

Q2. (i) For $t \neq 0$, X is Gaussian with zero mean (as B is), and continuous (again, as B is). The covariance of B is $\min(s, t)$. The covariance of X is

$$\begin{aligned}
 \text{cov}(X_s, X_t) &= \text{cov}(sB(1/s), tB(1/t)) \\
 &= E[sB(1/s).tB(1/t)] \\
 &= st.E[B(1/s)B(1/t)] \\
 &= st.\text{cov}(B(1/s), B(1/t)) \\
 &= st.\min(1/s, 1/t) = \min(t, s) = \min(s, t).
 \end{aligned}$$

This is the same covariance as Brownian motion. So, away from the origin, X is Brownian motion, as a Gaussian process is uniquely characterized by its mean and covariance (from the properties of the multivariate normal distribution). So X is continuous. So we can define it at the origin by continuity. So X is Brownian motion everywhere – X is BM.

(ii) Since Brownian motion is 0 at the origin, $X(0) = 0$. Since Brownian motion is continuous at the origin, $X(t) \rightarrow 0$ as $t \rightarrow 0$. This says that

$$tB(1/t) \rightarrow 0 \quad (t \rightarrow 0),$$

which is

$$B(t)/t \rightarrow 0 \quad (t \rightarrow \infty).$$

Note. For t integer, this is the Strong Law of Large Numbers applied to the distribution of $B(1)$, which is standard normal. The above neat proof by *time-inversion* follows from the proof of existence of Brownian motion (defined to be continuous), given in lectures by the wavelet expansion.

N. H. Bingham