smfsoln5.tex

SMF SOLUTIONS 5. 25.5.2012

Q1. (i) With X the number of successes, and as the prior in p is uniform,

$$P(X = x|p) = \binom{n}{x} p^x (1-p)^{n-x}, \qquad P(X = x) = \int_0^1 \binom{n}{x} p^x (1-p)^{n-x} dp,$$

 \mathbf{SO}

$$P(a
$$= \int_{a}^{b} {n \choose x} p^{x} (1-p)^{n-x} dp / B(x+1, n-x+1).$$$$

So the posterior is B(x+1, n-x+1). (ii) If the prior is now $B(\alpha, \beta)$, as in (i)

$$P(a
$$= \int_{a}^{b} {n \choose x} p^{x+\alpha-1} (1-p)^{n-x+\beta-1} dp.$$$$

So the posterior is $B(x + \alpha, n - x + \beta)$ (observe that U(0, 1) = B(1, 1), so (i) is the case $\alpha = \beta = 1$).

Q2. (i) For the Bernoulli distribution B(p), $f(x;p) = p^x (1-p)^{1-x}$,

$$\ell = x \log p + (1 - x) \log(1 - p), \qquad \ell' = \frac{x}{p} - \frac{1 - x}{1 - p},$$
$$\ell'' = -\frac{x}{p^2} - \frac{1 - x}{(1 - p)^2},$$
$$I(p) = -E[\ell''] = \frac{(1 - p)}{(1 - p)^2} + \frac{p}{p^2} = \frac{1}{p} + \frac{1}{1 - p} = \frac{1}{p(1 - p)}.$$

(ii) So the Jeffreys prior is $\pi(p) \propto \sqrt{I(p)} = 1/\sqrt{p(1-p)}$. This is the Beta distribution $B(\frac{1}{2}, \frac{1}{2})$, and $B(\frac{1}{2}, \frac{1}{2}) = \Gamma(\frac{1}{2}) \cdot \Gamma(\frac{1}{2})/\Gamma(1) = \pi$, as $\Gamma(\frac{1}{2}) = \sqrt{\pi}$. So the Jeffreys prior is

$$\pi(x) = \frac{1}{\pi\sqrt{x(1-x)}} \qquad (x \in [0,1]).$$

This is the *arc-sine law* (so called because the corresponding distribution function involves an arcsine). It is important in, e.g., fluctuations in cointossing. See e.g.

W. FELLER, An introduction to probability theory and its applications, Vol. 1, 3rd ed., Wiley, 1968, Ch. III.

Q3. Recall $\Gamma(z+1) = z\Gamma(z)$. $B(\alpha, \beta)$ has mean

$$E[X] = \int_0^1 x \cdot x^{\alpha - 1} (1 - x)^{\beta - 1} dx / B(\alpha, \beta) = \int_0^1 x^{\alpha} (1 - x)^{\beta - 1} dx / B(\alpha, \beta)$$
$$= B(\alpha + 1, \beta) / B(\alpha, \beta) = \frac{\Gamma(\alpha + 1)\Gamma(\beta)}{\Gamma(\alpha + \beta)} / \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)} \cdot = \alpha / (\alpha + \beta),$$

Q4. So the posterior mean in Q1(ii) is $(x + \alpha)/(n + \alpha + \beta)$. As the amount of data increases, $n \to \infty$, and by SLLN $x/n \to p$ a.s., where p is the true parameter value. With no data, x = n = 0, and the mean is the prior mean $\alpha/(\alpha + \beta)$. The value above is a compromise between these two.

Q5.

$$(f_{\alpha}*f_{\beta})(x) = \frac{1}{\Gamma(\alpha)\Gamma(\beta)} \int_{0}^{x} y^{\alpha-1} e^{-y} (x-y)^{\beta-1} e^{-(x-y)} dy = \frac{1}{\Gamma(\alpha)\Gamma(\beta)} e^{-x} \int_{0}^{x} y^{\alpha-1} (x-y)^{\beta-1} dy.$$

In the integral, I say, substitute y = xu. Then $I = x^{\alpha+\beta-1} \int_0^1 u^{\alpha-1} (1-u)^{\beta-1} du = x^{\alpha+\beta-1} B(\alpha,\beta)$. Combining, the RHS has the form of $f_{\alpha+\beta}(x)$ (to within constants!):

$$(f_{\alpha} * f_{\beta})(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \cdot B(\alpha, \beta) f_{\alpha + \beta}(x)$$

As both sides are densities, both integrate to 1. So the constant on the RHS is 1, which gives Euler's integral for the Beta function. NHB