smfsoln9.tex

## SMF SOLUTIONS 7. 31.5.2013

Q1. From the model equation

$$y_i = \sum_{j=1}^p a_{ij}\beta_j + \epsilon_i, \quad \epsilon_i \quad iid \quad N(0, \sigma^2),$$

the likelihood is

$$L = \frac{1}{\sigma^n 2\pi^{\frac{1}{2}n}} \cdot \prod_{i=1}^n \exp\{-\frac{1}{2}(y_i - \sum_{j=1}^p a_{ij}\beta_j)^2 / \sigma^2\}$$
  
=  $\frac{1}{\sigma^n 2\pi^{\frac{1}{2}n}} \cdot \exp\{-\frac{1}{2}\sum_{i=1}^n (y_i - \sum_{j=1}^p a_{ij}\beta_j)^2 / \sigma^2\},$ 

and the log-likelihood is

$$\ell := \log L = const - n \log \sigma - \frac{1}{2} \left[ \sum_{i=1}^{n} (y_i - \sum_{j=1}^{p} a_{ij} \beta_j)^2 \right] / \sigma^2.$$
 (\*)

We use Fisher's Method of Maximum, and maximise with respect to  $\beta_r$  in (\*) – or equivalently, the Method of Least Squares to minimise [...]:  $\partial \ell / \partial \beta_r = 0$  gives

$$\sum_{i=1}^{n} a_{ir}(y_i - \sum_{j=1}^{p} a_{ij}\beta_j) = 0 \qquad (r = 1, \dots, p),$$

or

$$\sum_{j=1}^{p} (\sum_{i=1}^{n} a_{ir} a_{ij}) \beta_j = \sum_{i=1}^{n} a_{ir} y_i.$$

Write  $C = (c_{ij})$  for the  $p \times p$  matrix

$$C := A^T A,$$

which we note is symmetric:  $C^T = C$ . Then

$$c_{ij} = \sum_{k=1}^{n} (A^T)_{ik} A_{kj} = \sum_{k=1}^{n} a_{ki} a_{kj}.$$

So this says

$$\sum_{j=1}^{p} c_{rj} \beta_j = \sum_{i=1}^{n} a_{ir} y_i = \sum_{i=1}^{n} (A^T)_{ri} y_i.$$

In matrix notation, this is

$$(C\beta)_r = (A^T y)_r \qquad (r = 1, \dots, p),$$

or combining,

$$C\beta = A^T y, \qquad C := A^T A. \tag{NE}$$

These are the normal equations. As A  $(n \times p$ , with  $p \ll n$  has full rank, A has rank p, so  $C := A^T A$  has rank p, so is non-singular. So the normal equations have solution

$$\hat{\beta} = C^{-1}A^T y = (A^T A)^{-1}A^T y.$$

Multiplying both sides by A,

$$Py = A(A^T A)^{-1} A^T y = A\hat{\beta}.$$

Q2.  $\partial \ell / \partial \sigma = 0$  gives  $-n/\sigma + [...]/\sigma^3 = 0$ ,  $\sigma^2 = [...]/n$ . At the maximum,  $\beta = \hat{\beta}$ , so [...] = SSE, giving  $\hat{\sigma}^2 = SSE/n$ .

$$SSE = (y - A\hat{\beta})^T (y - A\hat{\beta})$$
  
=  $y^T (I - P)^T (I - P) y$  (by Q1)  
=  $y^T (I - P) y$  ( $P^T = P, P^2 = P$  as P is a symmetric projection).

Q3. For an orthogonal matrix B,  $B^T B = I$ . So by the Product Theorem for Determinants,  $|B^T||B| = |I| = 1$ ,  $|B|^2 = 1$ , |B| = 1 (the case |B| = -1 corresponds to reversing the direction of an axis – anti-orthogonality – and we are about to take the modulus anyway). So the Jacobian of the transformation  $X \mapsto Y$  is mod det  $\partial y_i / \partial x_j = \text{mod det } b_{ij} = 1$  ( $y_i = \sum_j b_{ij} x_j$ ). Orthogonal transformations preserve lengths of vectors, so

$$||y|| = ||x|| : \sum y_i^2 = \sum x_i^2.$$

The joint density of the  $x_i$  is

$$f(x_1, \dots, x_n) = (2\pi)^{-\frac{1}{2}n} \exp\{-\frac{1}{2}\sum x_i^2\}.$$

So by the Jacobian formula, the joint density of the  $y_i$  is

$$f(y_1, \dots, y_n) = (2\pi)^{-\frac{1}{2}n} \exp\{-\frac{1}{2}\sum y_i^2\}$$

(for background on such uses of the Jacobian formula when changing variables in Statistics, see e.g. [BF], 2.2). This says that  $Y =_d X$ :  $y_1, \ldots, y_n$  are iid N(0, 1).

NHB