

smfsoln2(13a)

## SMF SOLUTIONS 2. 1.11.2013

Q1. (i) Consider

$$Q(\lambda) := \int (\lambda f + g)^2 = \lambda^2 \int f^2 + 2\lambda \int fg + \int g^2.$$

This is always non-negative, so its discriminant is  $\leq 0$  (if it were  $> 0$ , there would be two distinct real roots with a sign change between them). So (" $b^2 - 4ac \leq 0$ ")

$$\left(\int fg\right)^2 \leq \int f^2 \int g^2.$$

Equality holds iff  $Q$  has a double root,  $\lambda_0$  say. Then

$$Q(\lambda_0) = \int (\lambda_0 f + g)^2 = 0.$$

This forces

$$\lambda_0 f + g = 0$$

(a.e.), the required linear relation.

(ii) We apply this to  $X - E[X]$ ,  $Y - E[Y]$ , we find equality in CR iff a linear relation of the form  $a(X - E[X]) + b(Y - E[Y]) = 0$  holds. Taking  $X = \ell'$  (recall  $E[\ell'] = 0$ ),  $Y = u$  (recall  $EY = \theta$  as  $u$  is unbiased for  $\theta$ ), we find

$$u - \theta = a\ell'$$

for some  $a$  (so  $b = \theta$ ).

(iii) So

$$\ell' = u/a(\theta) - \theta/a(\theta) : \quad \ell = u \int d\theta/a(\theta) - \int \theta d\theta/a(\theta) + k(\mathbf{X}) :$$

$$L(\theta; \mathbf{X}) = \exp\{\alpha(\theta)u(\mathbf{X}) + \beta(\theta) + k(\mathbf{X})\}.$$

(iv) By Fisher-Neyman, this shows  $u$  is sufficient for  $\theta$ ; by Lehmann-Scheffé, this shows  $u$  is minimal sufficient for  $\theta$ .

Q2.

$$\ell = -n \log 2 - \sum |x_i - \theta|.$$

To maximise this – i.e. minimise  $\sum |x_i - \theta|$  – draw a graph. From this, the sum is minimised by  $\theta = Med$ , and increases linearly on either side of the sample median. So the MLE is  $\hat{\mu} = Med$ .

(ii) With one reading, as above,  $\ell$  decreases with slope -1 to the right of  $Med$ , slope +1 to the left of  $Med$ . So  $(\ell')^2 = 1$  (except at  $\lambda = Med$ , where the derivative is not defined – but we are going to integrate, and so can neglect null sets, let alone single points, so this does not matter). So  $I = \int (\partial \log f / \partial \theta)^2 f = \int f = 1$ , as  $f$  is a density. So the CR bound is  $1/n$ .

We are given that  $Med$  is asymptotically normal, and that its mean is  $med = \theta$ , so  $Med$  is asymptotically unbiased. By symmetry, the population median is  $med = \theta$ , where the density is  $\frac{1}{2}$ . So  $4f(med)^2 = 1$ , and the asymptotic variance of the sample median is  $1/n$ , the CR bound, so  $Med$  is also asymptotically efficient.

Q3. (i)

$$f(x; \mu) = \frac{1}{\pi(1 + (x - \mu)^2)}, \quad \ell = \log f = c - \log[1 + (x - \mu)^2],$$

$$\ell' = \frac{2(x - \mu)}{1 + (x - \mu)^2}, \quad \ell'(\mathbf{x}; \theta) = 2 \sum_1^n \frac{(x_i - \mu)}{1 + (x_i - \mu)^2}.$$

But (Q1) we have efficiency iff  $\ell'$  factorises in the form  $\ell'(\mathbf{x}; \theta) = A(\theta)(u(\mathbf{x}) - \theta)$ . The likelihood here does not factorise, so there is no efficient estimator.

(ii) The information per reading is

$$E[(\ell')^2] = \int (\partial f / \partial \mu)^2 f = \frac{4}{\pi} \int \frac{(x - \mu)^2}{[1 + (x - \mu)^2]^3} dx = \frac{4}{\pi} \int \frac{x^2}{[1 + x^2]^3} dx = \frac{4}{\pi} I,$$

say. One can evaluate  $I$  by Complex Analysis ( $f(z) := z^2/[1 + z^2]^3$ , round the contour  $\Gamma$  – semicircle in the upper half-plane on base  $[-R, R]$ ;  $f$  has a triple pole inside  $\Gamma$  of residue  $-i/16$ , so  $I = 2\pi i \text{ Res} = \pi/8$ ), giving the information per reading as  $\frac{1}{2}$ . So the information in a sample of size  $n$  is  $n/2$ , and the MLE has asymptotic variance  $2/n$ . As in Q2, the sample median has asymptotic variance  $\pi^2/4n$ . So the asymptotic efficiency is their ratio,  $8/\pi^2 \sim 81\%$ .  
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