

Solutions 9. 17.12.2010

Q1. (i) $EM(t) = 1$, as $N(0, 1)$ has MGF $e^{\frac{1}{2}s^2}$ and $N(0, t) =_d \sqrt{t}.N(0, 1)$.

$$\begin{aligned} E[M_t|\mathcal{F}_u] &= E[\exp\{s[B_u + (B_t - B_u)] - \tfrac{1}{2}ts^2\}|\mathcal{F}_u] \\ &= \exp\{sB_u - \tfrac{1}{2}us^2\}.E[\exp\{s(B_t - B_u) - \tfrac{1}{2}(t - u)s^2\}|\mathcal{F}_u], \end{aligned}$$

taking out what is known. The first term on RHS is M_u . Using the Strong Markov Property for BM to start afresh at time u , the second is $E[M_{t-u}]$, which is 1 by above. So M is a mg.

(ii) The stopping time T_n is bounded, so Doob's Stopping Time Principle gives $E[M(T_n)] = 1$:

$$\begin{aligned} 1 &= E \exp\{sB(T_n) - T_n \cdot \tfrac{1}{2}s^2\} \\ &= E[\exp\{sB(n) - n \cdot \tfrac{1}{2}s^2\}I(\tau > n)] + E[\exp\{sB(\tau_t) - \tau_t \cdot \tfrac{1}{2}s^2\}I(\tau \leq n)]. \end{aligned}$$

On $\tau > n$, $B(n) < t$, so the first term on RHS is at most

$$\exp\{st - n \cdot \tfrac{1}{2}s^2\}.P(\tau > n) \leq e^{st}.P(\tau > n) \rightarrow 0 \quad (n \rightarrow \infty).$$

Letting $n \rightarrow \infty$, $1 = E[\exp\{st - \tau_t \cdot \tfrac{1}{2}s^2\}.I(\tau_t < \infty)]$. But $\tau_t < \infty$ a.s. (otherwise B_u would lie below t for all u , so would have support bounded above; but each B_u is normal, so has support the whole line). So

$$1 = E[\exp\{st - \tau_t \cdot \tfrac{1}{2}s^2\}] : \quad E[\exp\{-\tau_t \cdot \tfrac{1}{2}s^2\}] = e^{-st}, \quad E[\exp\{-s\tau_t\}] = e^{-t\sqrt{2s}}.$$

(iii) The first-passage process τ is non-decreasing, as it takes longer to reach a higher level. Using the Strong Markov Property at time τ_t shows that the further time to first passage to level $t + u$ is independent of \mathcal{F}_t , and so of τ_t ; this says that the process τ has independent increments. This further time has the same distribution as τ_u , by the stationary-increments property of BM; so τ has stationary increments. so τ is a non-decreasing Lévy process, i.e. a subordinator.

(iv) By (iii), $E \exp\{-s\tau_{ct}\} = \exp\{-t \cdot c\sqrt{2s}\} = \exp\{-t \cdot \sqrt{2sc^2}\} = E \exp\{-sc^2\tau_t\}$.

Comparing, $c^2\tau_t = {}_d\tau_{ct}$: $\tau_t = {}_d\tau_{ct}/c^2$.

Q2.

$$\phi(s) = \frac{1}{\sqrt{2\pi}} \cdot \int_0^\infty \exp(-sx - \frac{1}{2x}) \cdot \frac{dx}{x^{3/2}}.$$

Differentiate under the integral sign (as we may, the integrand being monotone in s – we quote this):

$$\phi'(s) = -\frac{1}{\sqrt{2\pi}} \cdot \int_0^\infty \exp(-sx - \frac{1}{2x}) \cdot \frac{dx}{\sqrt{x}}.$$

The change of variable suggested interchanges the two terms in the exponential. It reverses the limits, and (check)

$$\frac{dx}{\sqrt{x}} = -\frac{1}{\sqrt{2s}} \cdot \frac{du}{u^{3/2}}.$$

This gives

$$\phi'(s) = -\frac{1}{\sqrt{2s}} \cdot \phi(s) : \quad \frac{\phi'(s)}{\phi(s)} = -\frac{1}{\sqrt{2s}}.$$

Integrate: $\log \phi(s) = -\sqrt{2s} + c$, $\phi(s) = ce^{-\sqrt{2s}}$. But $\phi(0) = \int f = 1$, so $\phi(s) = e^{-\sqrt{2s}}$. //

Q3. Adding independent random variables multiplies Laplace transforms (as with CFs – from the Multiplication Theorem), so $X_1 + \dots + X_n$ has Laplace transform $[\phi(s)]^n = e^{-n\sqrt{2s}}$. Replacing s by s/n^2 , $X_1 + \dots + X_n)/n^2$ has Laplace transform $\phi(s) = e^{-\sqrt{2s}}$, the Laplace transform of X . So $(X_1 + \dots + X_n)/n^2$ has the same distribution as X , as required.

This does not contradict the SLLN, as X has infinite mean.

NHB